

Bayesian Classification Example I

This live-script file simulates the Bayesian classification process within the context of a binary classification problem between classes ω_1 and ω_2 . We assume that the underlying feature space is one-dimensional and that the class-conditional probability distributions are given by the normal distribution parameterized by μ and σ^2 .

Let $P_1 = P(\omega_1)$ be the a-priori probability for the first class and $P_2 = P(\omega_2)$ be the a-priori probability for the second class. Moreover, the class-conditional probability distributions will be given as follows:

$$p(x|\omega_1; \mu_1, \sigma_1) = \frac{1}{\sqrt{2\pi} \sigma_1} e^{-\frac{(x-\mu_1)^2}{2\sigma_1^2}} \quad [1] \text{ and}$$

$$p(x|\omega_2; \mu_2, \sigma_2) = \frac{1}{\sqrt{2\pi} \sigma_2} e^{-\frac{(x-\mu_2)^2}{2\sigma_2^2}} \quad [2]$$

Bayesian classification will be conducted by identifying the corresponding decision regions in the one-dimensional feature space where the associated a-posteriori probabilities are maximized for each class. Thus, we may write:

$$R_1 = \{x \in \mathbb{R} : P(\omega_1|x) > P(\omega_2|x)\} \quad [3] \text{ and}$$

$$R_2 = \{x \in \mathbb{R} : P(\omega_2|x) > P(\omega_1|x)\} \quad [4]$$

```
% Clear workspace and command window.
clc
clear
% Initialize the prior probabilities for each class.
P1 = 0.5;
P2 = 1-P1;
% Initialize the internal parameters for the first Gaussian distribution.
mu_1 = 1;
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sigma_1 = 1;
% Initialize the internal parameters for the second Gaussian distribution.
mu_2 = -1;
sigma_2 = 1;

% Letting (mu_j,sigma_j) for j in {1,2} to denote the internal parameters
% of the Gaussian distributions for the two classes, we may define an extra
% variable k that parameterizes the interval in which each class-conditional
% probability density function is defined. In this context, we may define the
% interval for each class-conditional pdf as:
% X_j = [mu_j - k * sigma_j,mu_j + k * sigma_j] for j in {1,2}.
% In this context, the one-dimensional feature space X may be defined as:
% X = [mu_min - k * sigma_max,mu_max + k * sigma_max]
k = 3;
mu_min = min(mu_1,mu_2);
mu_max = max(mu_1,mu_2);
sigma_max = max(sigma_1,sigma_2);
% Set the dx parameter controlling the x-step.
dx = 0.001;
% Set the dy parameter controlling the y-step.
dy = 0.001;
% Define the one-dimensional feature space X.
X = mu_min-k*sigma_max:dx:mu_max+k*sigma_max;
% Compute the class-conditional probability distribution functions for the
% classes p_x_w1 and p_x_w2 within the previously defined interval.
p_x_w1 = normpdf(X,mu_1,sigma_1);
p_x_w2 = normpdf(X,mu_2,sigma_2);
% Compute the posterior probabilities for each class.
P_w1_x = p_x_w1 * P1;
P_w2_x = p_x_w2 * P2;
% Determine the minimum and maximum values of the X-axis.
Xmin = min(X);
Xmax = max(X);
% Determine the minimum and maximum values of the Y-axis.
Ymin = min([min(P_w1_x),min(P_w2_x)]);
Ymax = max([max(P_w1_x),max(P_w2_x)]);
% Define the Y-axis.
Y = Ymin:dy:Ymax;
% Plot the posterior probabilities for each class.
figure('Name','Bayesian Classification: Posterior Probabilities');
title('Bayesian Classification: Posterior Class Probabilities');
hold on
plot(X,P_w1_x,'-r','LineWidth',2.5);
plot(X,P_w2_x,'-b','LineWidth',2.5);
xlabel('x');
ylabel('P(w|x)');
% Shade the area below each probability density function.
H1=area(X,P_w1_x,'FaceColor','r');
H1.FaceAlpha = 0.2;
H2=area(X,P_w2_x,'FaceColor','b');
H2.FaceAlpha = 0.2;

```

In this framework, the Bayesian classification rule may be formulated by the following discrimination function:

$$g(x) = \begin{cases} \omega_1, & x \in R_1 \\ \omega_2, & x \in R_2 \end{cases} \quad [5]$$

Therefore, the decision boundary

$R_0 = \{x \in \mathbb{R} : h(x) = P(\omega_1|x) - P(\omega_2|x) = 0\}$ [6] between the classification regions R_1 and R_2 may be computed by solving the following equation:

$$P(\omega_1|x) = P(\omega_2|x) \Rightarrow$$

$$\frac{p(x|\omega_1)P(\omega_1)}{p(x)} = \frac{p(x|\omega_2)P(\omega_2)}{p(x)} \Rightarrow$$

$$\frac{p(x|\omega_1)}{p(x|\omega_2)} = \frac{P(\omega_2)}{P(\omega_1)} \Rightarrow$$

$$e^{\frac{(x-\mu_2)^2}{2\sigma_2^2} - \frac{(x-\mu_1)^2}{2\sigma_1^2}} = \frac{P(\omega_2)}{P(\omega_1)} \frac{\sigma_1}{\sigma_2} \quad [7]$$

Letting $z = \frac{(x - \mu_2)^2}{2\sigma_2^2} - \frac{(x - \mu_1)^2}{2\sigma_1^2}$ [8], we may write that

$$z = \ln \frac{P(\omega_2)}{P(\omega_1)} + \ln \frac{\sigma_1}{\sigma_2} \quad [9]$$

Moreover, by letting $\lambda_p = \ln \frac{P(\omega_2)}{P(\omega_1)}$ [10] and $\lambda_\sigma = \ln \frac{\sigma_1}{\sigma_2}$ [11], Equation (6) may be re-written as:

$$\frac{(x - \mu_2)^2}{2\sigma_2^2} - \frac{(x - \mu_1)^2}{2\sigma_1^2} = \lambda_p + \lambda_\sigma \quad [12]$$

Equation (12) suggests that the decision boundary between classes ω_1 and ω_2 will be given by the solutions of the following equation:

$$h(x) = (\sigma_1^2 - \sigma_2^2)x^2 + 2(\mu_1\sigma_2^2 - \mu_2\sigma_1^2)x + \sigma_1^2\mu_2^2 - \sigma_2^2\mu_1^2 - 2\sigma_1^2\sigma_2^2(\lambda_p + \lambda_\sigma) = 0 \quad [13]$$

By letting $\alpha = (\sigma_1^2 - \sigma_2^2)$ [14], $\beta = 2(\mu_1\sigma_2^2 - \mu_2\sigma_1^2)$ [15]

and $\gamma = \sigma_1^2\mu_2^2 - \sigma_2^2\mu_1^2 - 2\sigma_1^2\sigma_2^2(\lambda_p + \lambda_\sigma)$ [16], we may write that

$h(x) = \alpha x^2 + \beta x + \gamma$ [17]. Thus, the discriminant for the second degree polynomial $h(x)$ can be expressed as: $\Delta_h = \beta^2 - 4\alpha\gamma$.

In this context, we may identify the following cases for the decision boundary $R_0 = \{x \in \mathbb{R} : h(x) = 0\}$ [18].

$$\text{Case I: } \alpha = 0, \beta \neq 0 \Rightarrow R_0 = \{x_0\} = \left\{ -\frac{\gamma}{\beta} \right\} \quad (h(x) = \beta x + \gamma)$$

$$\text{Case II: } \alpha \neq 0, \Delta_h > 0 \Rightarrow R_0 = \{x_0^+, x_0^-\} = \left\{ \frac{-\beta + \sqrt{\Delta_h}}{2\alpha}, \frac{-\beta - \sqrt{\Delta_h}}{2\alpha} \right\}$$

$$(h(x) = \alpha(x - x_0^-)(x - x_0^+))$$

In fact, we set that $x_0^+ = \max \left\{ \frac{-\beta + \sqrt{\Delta_h}}{2\alpha}, \frac{-\beta - \sqrt{\Delta_h}}{2\alpha} \right\}$ and

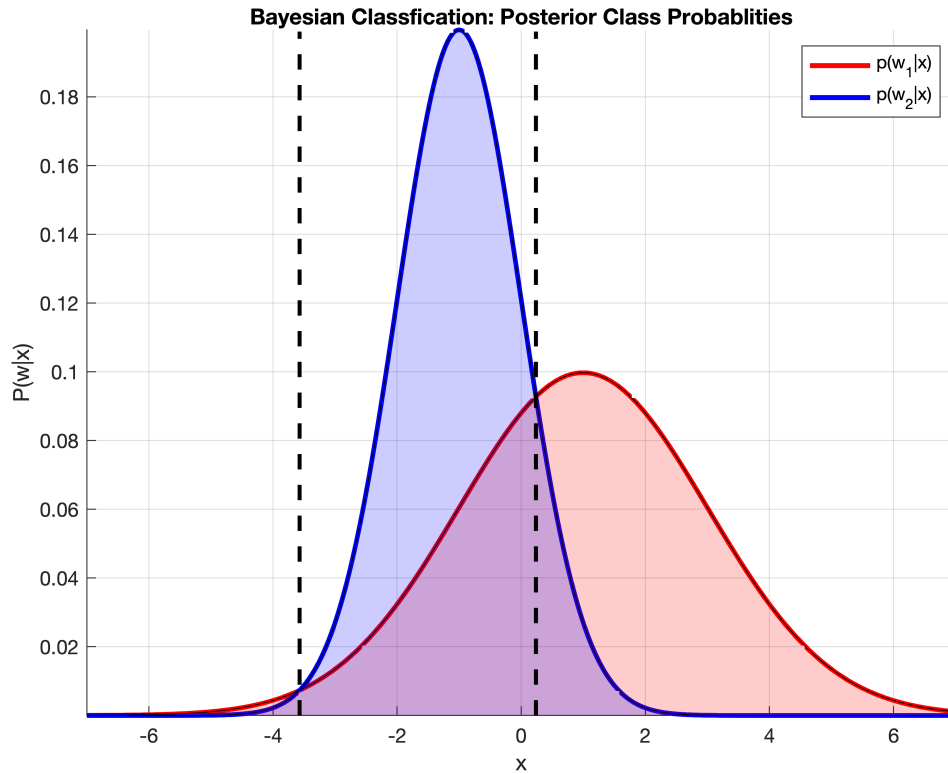
$$x_0^- = \min \left\{ \frac{-\beta + \sqrt{\Delta_h}}{2\alpha}, \frac{-\beta - \sqrt{\Delta_h}}{2\alpha} \right\}$$

Case III: $\alpha \neq 0, \Delta_h = 0 \Rightarrow R_0 = \{\hat{x}_0\} = \left\{ -\frac{\beta}{2\alpha} \right\}$ ($h(x) = \alpha(x - \hat{x}_0)^2$)

Case IV: $\alpha \neq 0, \Delta_h < 0 \Rightarrow R_0 = \{\emptyset\}$

Cases III and IV will not be considered since they involve occasions where we have either $h(x) \geq 0, \forall x \in \mathbb{R}$ or $h(x) \leq 0, \forall x \in \mathbb{R}$ respectively, which, in turn, result in the following degenerate cases $R_2 = \{\emptyset\}$ or $R_1 = \{\emptyset\}$.

```
% Compute the constant values Lp and Ls.
Lp = log(P2/P1);
Ls = log(sigma_1/sigma_2);
% Compute the coefficients of the polynomial h(x) = A*x^2 + B*x + C whose roots
% determine the boundary region Ro between the two classes.
A = sigma_1^2 - sigma_2^2;
B = 2*(mu_1*sigma_2^2 - mu_2*sigma_1^2);
C = (sigma_1*mu_2)^2 - (sigma_2*mu_1)^2 - 2*((sigma_1*sigma_2)^2)*(Lp+Ls);
% Compute the discriminant D of the polynomial h(x).
D = B^2 - 4*A*C;
% Compute the solution points for the quadratic equation: h(x) = 0 and plot
% the corresponding boundary.
if(A==0)
    Xo = - C/B;
    plot(Xo*ones(1,length(Y)),Y,'--k','LineWidth',2.0);
else
    if(D>0)
        Xo_plus = max((-B + sqrt(D)) / (2*A), (-B - sqrt(D)) / (2*A));
        Xo_minus = min((-B + sqrt(D)) / (2*A), (-B - sqrt(D)) / (2*A));
        plot(Xo_plus*ones(1,length(Y)),Y,'--k','LineWidth',2.0);
        plot(Xo_minus*ones(1,length(Y)),Y,'--k','LineWidth',2.0);
    end
end
hold off
legend({'p(w_1|x)', 'p(w_2|x)'});
grid on
axis([Xmin Xmax Ymin Ymax]);
```



According to the previous analysis, decision regions R_1 and R_2 may be reformulated as: $R_1 = \{x \in \mathbb{R} : h(x) > 0\}$ [19] and $R_2 = \{x \in \mathbb{R} : h(x) < 0\}$ [20]

Thus, we may consider the following cases for the decision regions R_1 and R_2 :

Case I: $\alpha = 0, \beta \neq 0 \Rightarrow R_1 = (x_0, +\infty)$ and $R_2 = (-\infty, x_0)$

Case IIa: $\alpha > 0, \Delta_h > 0 \Rightarrow R_1 = (-\infty, x_0^-) \cup (x_0^+, +\infty)$ and $R_2 = (x_0^-, x_0^+)$

Case IIb: $\alpha < 0, \Delta_h > 0 \Rightarrow R_1 = (x_0^-, x_0^+)$ and $R_2 = (-\infty, x_0^-) \cup (x_0^+, +\infty)$

In this setting, the overall probability of error is given by the following formula:

$$P(\text{error}) = P(\omega_1) \int_{R_2} p(x|\omega_1)dx + P(\omega_2) \int_{R_1} p(x|\omega_2)dx \quad [22]$$

Case I: Equation (22) may be rewritten in the following form:

$$P(error) = P(\omega_1) \int_{-\infty}^{x_0} p(x|\omega_1)dx + P(\omega_2) \int_{x_0}^{+\infty} p(x|\omega_2)dx \quad [23]$$

Case IIa: Equation (22) may be rewritten in the following form:

$$P(error) = P(\omega_1) \int_{x_0^-}^{x_0^+} p(x|\omega_1)dx + P(\omega_2) \left[\int_{-\infty}^{x_0^-} p(x|\omega_2)dx + \int_{x_0^+}^{+\infty} p(x|\omega_2)dx \right] \quad [24]$$

Case IIb: Equation (22) may be rewritten in the following form:

$$P(error) = P(\omega_1) \left[\int_{-\infty}^{x_0^-} p(x|\omega_1)dx + \int_{x_0^+}^{+\infty} p(x|\omega_1)dx \right] + P(\omega_2) \int_{x_0^-}^{x_0^+} p(x|\omega_2)dx \quad [25]$$

Evaluating the integrals appearing within Equations (23), (24) and (25) requires the existence of a computational procedure for estimating the cumulative distribution function that corresponds to a given probability distribution function. For the case of a normal probability distribution function, the associated cumulative distribution function is given by:

$$\int_{-\infty}^x \frac{1}{\sqrt{2\pi} \sigma} e^{-\frac{(t-\mu)^2}{2\sigma^2}} dt = I(x; \mu, \sigma) = \text{normcdf}(x, \mu, \sigma) \quad [26]$$

Therefore, we may define the following auxiliary integrals as:

$$\int_x^{\infty} \frac{1}{\sqrt{2\pi} \sigma} e^{-\frac{(t-\mu)^2}{2\sigma^2}} dt = 1 - I(x; \mu, \sigma) \quad [27]$$

and

$$\int_{x_a}^{x_b} \frac{1}{\sqrt{2\pi} \sigma} e^{-\frac{(t-\mu)^2}{2\sigma^2}} dt = I(x_b; \mu, \sigma) - I(x_a; \mu, \sigma) \quad [28]$$

In view of Equations (26), (27) and (28) the overall error probabilities for the previously identified cases may be given as:

Case I: $P(\text{error}) = P(\omega_1)I(x_0; \mu_1, \sigma_1) + P(\omega_2)[1 - I(x_0; \mu_2, \sigma_2)]$ [29]

Case IIa:

$P(\text{error}) = P(\omega_1)[I(x_0^+; \mu_1, \sigma_1) - I(x_0^-; \mu_1, \sigma_1)] + P(\omega_2)[I(x_0^-; \mu_2, \sigma_2) + 1 - I(x_0^+; \mu_2, \sigma_2)]$ [30]

Case IIb: $P(\text{error}) =$

$P(\omega_1)[I(x_0^-; \mu_1, \sigma_1) + 1 - I(x_0^+; \mu_1, \sigma_1)] + P(\omega_2)[I(x_0^+; \mu_2, \sigma_2) - I(x_0^-; \mu_2, \sigma_2)]$ [31]

```
% Compute the overall error probabilities for each one of the previously
% identified cases.
if(A==0)
    Pe = P1 * normcdf(Xo,mu_1,sigma_1) + P2 * (1 - normcdf(Xo,mu_2,sigma_2))
else
    if (D > 0)
        if (A > 0)
            Pe = P1 * (normcdf(Xo_plus,mu_1,sigma_1) - normcdf(Xo_minus,mu_1,sigma_1))
        else
            Pe = P1 * (normcdf(Xo_minus,mu_1,sigma_1) + 1 - normcdf(Xo_plus,mu_1,sigma_1))
        end
    end
end
end
```

Pe = 0.2267